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#### Allerton Conference 2018

A. Makur (MIT)

Capacity of BSC Permutation Channel

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## Outline

### Introduction

- Motivation: Coding for Communication Networks
- The Permutation Channel Model
- Capacity of the BSC Permutation Channel

## 2 Achievability

## 3 Converse

## 4 Conclusion



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#### Model communication network as a channel:

• Alphabet symbols = all possible *L*-bit packets  $\Rightarrow 2^{L}$  input symbols



- Alphabet symbols = all possible *L*-bit packets
- multipath routed network or evolving network topology



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Consider a communication network where packets can be dropped:



- *n*-length codeword = sequence of *n* packets
- Equivalent Erasure channel: Erase each symbol/packet of codeword independently with probability p ∈ (0, 1)
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- *n*-length codeword = sequence of *n* packets
- Erasure channel: Erase each symbol/packet of codeword independently with probability p ∈ (0, 1)
- Random permutation block: Randomly permute packets of codeword
- Coding: Add sequence numbers (packet size =  $L + \log(n)$  bits, alphabet size =  $n2^{L}$ )

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- Random permutation block: Randomly permute packets of codeword
- Coding: Add sequence numbers and use standard coding techniques
- More refined coding techniques *simulate* sequence numbers,
  - e.g. [Mitzenmacher 2006], [Metzner 2009]

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How do you code in such channels without increasing alphabet size?

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- Discrete memoryless channel  $P_{Z|X}$  with input and output alphabets  $\mathcal{X}$  and  $\mathcal{Y}$  produces  $Z_1^n$ :

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- Possibly randomized decoder  $g_n : \mathcal{Y}^n \to \mathcal{M}$  produces estimate  $\hat{\mathcal{M}} = g_n(Y_1^n)$  at receiver



#### General Principle:

"Encode the information in an object that is invariant under the [permutation] transformation." [Kovačević-Vukobratović 2013]



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# What about the information theoretic aspects of this model?



• Average probability of error  $P_{\text{error}}^n \triangleq \mathbb{P}(M \neq \hat{M})$ 



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#### Definition (Permutation Channel Capacity)

$$C_{\mathsf{perm}}(P_{Z|X}) \triangleq \sup\{R \ge 0 : R \text{ is achievable}\}$$

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• Channel is binary symmetric channel, denoted BSC(*p*):

$$\forall z, x \in \{0, 1\}, \ P_{Z|X}(z|x) = \begin{cases} 1-p, & \text{for } z = x \\ p, & \text{for } z \neq x \end{cases}$$



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#### Main Question

What is the permutation channel capacity of the BSC?

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### Introduction

#### 2 Achievability

- Encoder and Decoder
- Testing between Converging Hypotheses
- Intuition via Central Limit Theorem
- Second Moment Method for TV Distance

### 3 Converse





• Fix a message  $m \in \{0, 1\}$ 



0 
$$q_0 = \frac{1}{3}$$
  $q_1 = \frac{2}{3}$  1



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• Memoryless BSC(p) outputs  $Z_1^n \stackrel{\text{i.i.d.}}{\sim} \text{Ber}(p * q_m)$ , where  $p * q_m \triangleq p(1 - q_m) + q_m(1 - p)$  is the convolution of p and  $q_m$ 



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• Maximum Likelihood (ML) decoder:  $\hat{M} = \mathbb{1}\left\{\frac{1}{n}\sum_{i=1}^{n}Y_{i} \geq \frac{1}{2}\right\}$ 



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$$\frac{1}{n} \sum_{i=1}^{n} Y_i \rightarrow p * q_m$$
 in probability as  $n \rightarrow \infty$  [WLLN]  
 $\Rightarrow \lim_{n \rightarrow \infty} P_{\text{error}}^n = 0$  as  $p * q_0 \neq p * q_1$ 

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- Challenge: Although <sup>1</sup>/<sub>n</sub> ∑<sup>n</sup><sub>i=1</sub> Y<sub>i</sub> → p \* <sup>m</sup>/<sub>n<sup>R</sup></sub> in probability as n → ∞, consecutive messages become indistinguishable i.e. <sup>m</sup>/<sub>n<sup>R</sup></sub> <sup>m+1</sup>/<sub>n<sup>R</sup></sub> → 0

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# What is the largest *R* such that two consecutive messages can be distinguished?

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Capacity of BSC Permutation Channel

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Given 
$$H = 0: X_1^n \stackrel{\text{i.i.d.}}{\sim} P_{X|H=0} = \text{Ber}(q)$$
  
Given  $H = 1: X_1^n \stackrel{\text{i.i.d.}}{\sim} P_{X|H=1} = \text{Ber}\left(q + \frac{1}{n^R}\right)$ 

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$$T_n \triangleq \frac{1}{n} \sum_{i=1}^n X_i - q - \frac{1}{2n^R}$$

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• Let  $\hat{H}_{ML}^{n}(T_{n})$  denote the ML decoder for H based on  $T_{n}$  with minimum probability of error  $P_{ML}^{n} \triangleq \mathbb{P}(\hat{H}_{ML}^{n}(T_{n}) \neq H)$ 

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- Want: Largest R > 0 such that  $\lim_{n \to \infty} P_{ML}^n = 0$ ?

• For large *n*,  $P_{T_n|H}(\cdot|0)$  and  $P_{T_n|H}(\cdot|1)$  are Gaussian distributions [CLT]

Figure:



For large n, P<sub>Tn|H</sub>(·|0) and P<sub>Tn|H</sub>(·|1) are Gaussian distributions [CLT]
 |ℝ[Tn|H = 0] - ℝ[Tn|H = 1]| = 1/n<sup>R</sup>

Figure:



- For large *n*,  $P_{T_n|H}(\cdot|0)$  and  $P_{T_n|H}(\cdot|1)$  are Gaussian distributions [CLT]
- $|\mathbb{E}[T_n|H=0] \mathbb{E}[T_n|H=1]| = 1/n^R$
- Standard deviations are  $\Theta(1/\sqrt{n})$

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$$|\mathbb{E}[T_n|H=0] - \mathbb{E}[T_n|H=1]| = 1/n^{t}$$

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- **Case**  $R > \frac{1}{2}$ : Decoding is impossible  $\otimes$



Lemma (2<sup>nd</sup> Moment Method [Evans-Kenyon-Peres-Schulman 2000])

$$\left\| P_{\mathcal{T}_n | \mathcal{H}=1} - P_{\mathcal{T}_n | \mathcal{H}=0} \right\|_{\mathsf{TV}} \geq \frac{\left( \mathbb{E}[\mathcal{T}_n | \mathcal{H}=1] - \mathbb{E}[\mathcal{T}_n | \mathcal{H}=0] \right)^2}{4 \, \mathbb{VAR}(\mathcal{T}_n)}$$

where  $||P - Q||_{TV} = \frac{1}{2} ||P - Q||_{\ell^1}$  is the total variation (TV) distance between the distributions P and Q.

Lemma (2<sup>nd</sup> Moment Method [Evans-Kenyon-Peres-Schulman 2000])

$$\left\| \boldsymbol{P}_{T_n|H=1} - \boldsymbol{P}_{T_n|H=0} \right\|_{\mathsf{TV}} \geq \frac{\left( \mathbb{E}[T_n|H=1] - \mathbb{E}[T_n|H=0] \right)^2}{4 \, \mathbb{VAR}(T_n)}$$

where  $||P - Q||_{TV} = \frac{1}{2} ||P - Q||_{\ell^1}$  is the total variation (TV) distance between the distributions P and Q.

**Proof:** Let 
$$T_n^+ \sim P_{T_n|H=1}$$
 and  $T_n^- \sim P_{T_n|H=0}$   
 $\left(\mathbb{E}[T_n^+] - \mathbb{E}[T_n^-]\right)^2 = \left(\sum_t t \left(P_{T_n|H}(t|1) - P_{T_n|H}(t|0)\right)\right)^2$ 

Lemma (2<sup>nd</sup> Moment Method [Evans-Kenyon-Peres-Schulman 2000])

$$\left|P_{T_n|H=1} - P_{T_n|H=0}\right|_{\mathsf{TV}} \geq \frac{\left(\mathbb{E}[T_n|H=1] - \mathbb{E}[T_n|H=0]\right)^2}{4 \, \mathbb{VAR}(T_n)}$$

where  $||P - Q||_{TV} = \frac{1}{2} ||P - Q||_{\ell^1}$  is the total variation (TV) distance between the distributions P and Q.

**Proof:** Let 
$$T_n^+ \sim P_{T_n|H=1}$$
 and  $T_n^- \sim P_{T_n|H=0}$   
 $\left(\mathbb{E}[T_n^+] - \mathbb{E}[T_n^-]\right)^2 = \left(\sum_t t \sqrt{P_{T_n}(t)} \frac{\left(P_{T_n|H}(t|1) - P_{T_n|H}(t|0)\right)}{\sqrt{P_{T_n}(t)}}\right)^2$ 

Lemma (2<sup>nd</sup> Moment Method [Evans-Kenyon-Peres-Schulman 2000])

$$\left\| \boldsymbol{P}_{T_n|H=1} - \boldsymbol{P}_{T_n|H=0} \right\|_{\mathsf{TV}} \geq \frac{\left( \mathbb{E}[T_n|H=1] - \mathbb{E}[T_n|H=0] \right)^2}{4 \, \mathbb{VAR}(T_n)}$$

where  $||P - Q||_{TV} = \frac{1}{2} ||P - Q||_{\ell^1}$  is the total variation (TV) distance between the distributions P and Q.

#### Proof: Cauchy-Schwarz inequality

$$\left( \mathbb{E} \left[ T_n^+ \right] - \mathbb{E} \left[ T_n^- \right] \right)^2 = \left( \sum_t t \sqrt{P_{T_n}(t)} \frac{\left( P_{T_n \mid \mathcal{H}}(t \mid 1) - P_{T_n \mid \mathcal{H}}(t \mid 0) \right)}{\sqrt{P_{T_n}(t)}} \right)^2 \\ \leq \left( \sum_t t^2 P_{T_n}(t) \right) \left( \sum_t \frac{\left( P_{T_n \mid \mathcal{H}}(t \mid 1) - P_{T_n \mid \mathcal{H}}(t \mid 0) \right)^2}{P_{T_n}(t)} \right)$$

Lemma (2<sup>nd</sup> Moment Method [Evans-Kenyon-Peres-Schulman 2000])

$$\left| P_{T_n|H=1} - P_{T_n|H=0} \right|_{\mathsf{TV}} \geq \frac{\left( \mathbb{E}[T_n|H=1] - \mathbb{E}[T_n|H=0] \right)^2}{4 \, \mathbb{VAR}(T_n)}$$

where  $||P - Q||_{TV} = \frac{1}{2} ||P - Q||_{\ell^1}$  is the total variation (TV) distance between the distributions P and Q.

**Proof:** Recall that  $T_n$  is zero-mean

$$\left( \mathbb{E} \left[ T_n^+ \right] - \mathbb{E} \left[ T_n^- \right] \right)^2 = \left( \sum_t t \sqrt{P_{T_n}(t)} \frac{\left( P_{T_n|H}(t|1) - P_{T_n|H}(t|0) \right)}{\sqrt{P_{T_n}(t)}} \right)^2 \\ \leq \mathbb{VAR}(T_n) \left( \sum_t \frac{\left( P_{T_n|H}(t|1) - P_{T_n|H}(t|0) \right)^2}{P_{T_n}(t)} \right)$$

Lemma (2<sup>nd</sup> Moment Method [Evans-Kenyon-Peres-Schulman 2000])

$$\left|P_{T_n|H=1} - P_{T_n|H=0}\right|_{\mathsf{TV}} \geq \frac{\left(\mathbb{E}[T_n|H=1] - \mathbb{E}[T_n|H=0]\right)^2}{4 \, \mathbb{VAR}(T_n)}$$

where  $||P - Q||_{TV} = \frac{1}{2} ||P - Q||_{\ell^1}$  is the total variation (TV) distance between the distributions P and Q.

Proof: Hammersley-Chapman-Robbins bound

$$\left(\mathbb{E}[T_n^+] - \mathbb{E}[T_n^-]\right)^2 = \left(\sum_t t \sqrt{P_{\mathcal{T}_n}(t)} \frac{\left(P_{\mathcal{T}_n|H}(t|1) - P_{\mathcal{T}_n|H}(t|0)\right)}{\sqrt{P_{\mathcal{T}_n}(t)}}\right)^2$$
$$\leq 4 \operatorname{VAR}(\mathcal{T}_n) \underbrace{\left(\frac{1}{4}\sum_t \frac{\left(P_{\mathcal{T}_n|H}(t|1) - P_{\mathcal{T}_n|H}(t|0)\right)^2}{P_{\mathcal{T}_n}(t)}\right)}_{H_{\mathcal{T}_n}(t)}$$

Vincze-Le Cam distance

Lemma (2<sup>nd</sup> Moment Method [Evans-Kenyon-Peres-Schulman 2000])

$$\left\| \boldsymbol{P}_{T_n|H=1} - \boldsymbol{P}_{T_n|H=0} \right\|_{\mathsf{TV}} \geq \frac{\left( \mathbb{E}[T_n|H=1] - \mathbb{E}[T_n|H=0] \right)^2}{4 \, \mathbb{VAR}(T_n)}$$

where  $||P - Q||_{TV} = \frac{1}{2} ||P - Q||_{\ell^1}$  is the total variation (TV) distance between the distributions P and Q.

Proof:

$$\begin{split} \left( \mathbb{E} \left[ T_n^+ \right] - \mathbb{E} \left[ T_n^- \right] \right)^2 &= \left( \sum_t t \sqrt{P_{\mathcal{T}_n}(t)} \, \frac{\left( P_{\mathcal{T}_n|H}(t|1) - P_{\mathcal{T}_n|H}(t|0) \right)}{\sqrt{P_{\mathcal{T}_n}(t)}} \right)^2 \\ &\leq 4 \, \mathbb{VAR}(\mathcal{T}_n) \left( \frac{1}{4} \sum_t \frac{\left( P_{\mathcal{T}_n|H}(t|1) - P_{\mathcal{T}_n|H}(t|0) \right)^2}{P_{\mathcal{T}_n}(t)} \right) \\ &\leq 4 \, \mathbb{VAR}(\mathcal{T}_n) \left\| P_{\mathcal{T}_n|H=1} - P_{\mathcal{T}_n|H=0} \right\|_{\mathsf{TV}} \end{split}$$

# Achievability Proof

#### Theorem (Achievability)

For any 0 < R < 1/2, consider the binary hypothesis testing problem with  $H \sim \text{Ber}(\frac{1}{2})$ , and  $X_1^n \stackrel{\text{i.i.d.}}{\sim} \text{Ber}(q + \frac{h}{n^R})$  given  $H = h \in \{0, 1\}$ .

Proof: Start with Le Cam's relation

$$P_{ML}^{n} = \frac{1}{2} \left( 1 - \left\| P_{T_{n}|H=1} - P_{T_{n}|H=0} \right\|_{TV} \right)$$
#### Theorem (Achievability)

For any 0 < R < 1/2, consider the binary hypothesis testing problem with  $H \sim \text{Ber}(\frac{1}{2})$ , and  $X_1^n \stackrel{\text{i.i.d.}}{\sim} \text{Ber}(q + \frac{h}{n^R})$  given  $H = h \in \{0, 1\}$ .

#### Proof: Apply second moment method lemma

$$\begin{aligned} P_{\mathsf{ML}}^{n} &= \frac{1}{2} \left( 1 - \left\| P_{\mathcal{T}_{n}|\mathcal{H}=1} - P_{\mathcal{T}_{n}|\mathcal{H}=0} \right\|_{\mathsf{TV}} \right) \\ &\leq \frac{1}{2} \left( 1 - \frac{\left( \mathbb{E}[\mathcal{T}_{n}|\mathcal{H}=1] - \mathbb{E}[\mathcal{T}_{n}|\mathcal{H}=0] \right)^{2}}{4 \, \mathbb{VAR}(\mathcal{T}_{n})} \right) \end{aligned}$$

#### Theorem (Achievability)

For any 0 < R < 1/2, consider the binary hypothesis testing problem with  $H \sim \text{Ber}(\frac{1}{2})$ , and  $X_1^n \stackrel{\text{i.i.d.}}{\sim} \text{Ber}(q + \frac{h}{n^R})$  given  $H = h \in \{0, 1\}$ .

Proof: After explicit computation and simplification...

$$egin{split} P_{\mathsf{ML}}^n &= rac{1}{2} \left( 1 - \left\| P_{\mathcal{T}_n \mid H=1} - P_{\mathcal{T}_n \mid H=0} 
ight\|_{\mathsf{TV}} 
ight) \ &\leq rac{1}{2} \left( 1 - rac{\left( \mathbb{E}[\mathcal{T}_n \mid H=1] - \mathbb{E}[\mathcal{T}_n \mid H=0] 
ight)^2}{4 \, \mathbb{VAR}(\mathcal{T}_n)} 
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**Proof:** For any  $0 < R < \frac{1}{2}$ ,

$$P_{\mathsf{ML}}^{n} = \frac{1}{2} \left( 1 - \left\| P_{\mathcal{T}_{n}|\mathcal{H}=1} - P_{\mathcal{T}_{n}|\mathcal{H}=0} \right\|_{\mathsf{TV}} \right)$$
$$\leq \frac{1}{2} \left( 1 - \frac{\left( \mathbb{E}[\mathcal{T}_{n}|\mathcal{H}=1] - \mathbb{E}[\mathcal{T}_{n}|\mathcal{H}=0] \right)^{2}}{4 \, \mathbb{VAR}(\mathcal{T}_{n})} \right)$$
$$\leq \frac{3}{2n^{1-2R}}$$

#### Theorem (Achievability)

For any 0 < R < 1/2, consider the binary hypothesis testing problem with  $H \sim \text{Ber}(\frac{1}{2})$ , and  $X_1^n \stackrel{\text{i.i.d.}}{\sim} \text{Ber}(q + \frac{h}{n^R})$  given  $H = h \in \{0, 1\}$ . Then,  $\lim_{n \to \infty} P_{\text{ML}}^n = 0$ .

**Proof:** For any  $0 < R < \frac{1}{2}$ ,

$$P_{\mathsf{ML}}^{n} = \frac{1}{2} \left( 1 - \left\| P_{\mathcal{T}_{n}|\mathcal{H}=1} - P_{\mathcal{T}_{n}|\mathcal{H}=0} \right\|_{\mathsf{TV}} \right)$$
  
$$\leq \frac{1}{2} \left( 1 - \frac{\left( \mathbb{E}[\mathcal{T}_{n}|\mathcal{H}=1] - \mathbb{E}[\mathcal{T}_{n}|\mathcal{H}=0] \right)^{2}}{4 \, \mathbb{VAR}(\mathcal{T}_{n})} \right)$$
  
$$\leq \frac{3}{2n^{1-2R}} \to 0 \text{ as } n \to \infty$$

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For any 0 < R < 1/2, consider the binary hypothesis testing problem with  $H \sim \text{Ber}(\frac{1}{2})$ , and  $X_1^n \stackrel{\text{i.i.d.}}{\sim} \text{Ber}(q + \frac{h}{n^R})$  given  $H = h \in \{0, 1\}$ . Then,  $\lim_{n \to \infty} P_{\text{ML}}^n = 0$ . This implies that:  $C_{\text{perm}}(\text{BSC}(p)) \ge \frac{1}{2}$ .

**Proof:** For any  $0 < R < \frac{1}{2}$ ,

$$P_{\mathsf{ML}}^{n} = \frac{1}{2} \left( 1 - \left\| P_{\mathcal{T}_{n}|\mathcal{H}=1} - P_{\mathcal{T}_{n}|\mathcal{H}=0} \right\|_{\mathsf{TV}} \right)$$
$$\leq \frac{1}{2} \left( 1 - \frac{\left( \mathbb{E}[\mathcal{T}_{n}|\mathcal{H}=1] - \mathbb{E}[\mathcal{T}_{n}|\mathcal{H}=0] \right)^{2}}{4 \, \mathbb{VAR}(\mathcal{T}_{n})} \right)$$
$$\leq \frac{3}{2n^{1-2R}} \to 0 \text{ as } n \to \infty$$

#### Introduction

#### 2 Achievability

#### 3 Converse

- Fano's Inequality Argument
- CLT Approximation

### Conclusion

• Consider the Markov chain  $M \to X_1^n \to Z_1^n \to Y_1^n \to S_n \triangleq \sum_{i=1}^n Y_i$ , and a sequence of encoder-decoder pairs  $\{(f_n, g_n)\}_{n \in \mathbb{N}}$  such that  $|\mathcal{M}| = n^R$  and  $\lim_{n \to \infty} P_{\text{error}}^n = 0$ 

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- Standard argument, cf. [Cover-Thomas 2006]: *M* is uniform

 $R\log(n) = H(M)$ 

- Consider the Markov chain  $M \to X_1^n \to Z_1^n \to Y_1^n \to S_n \triangleq \sum_{i=1}^n Y_i$ , and a sequence of encoder-decoder pairs  $\{(f_n, g_n)\}_{n \in \mathbb{N}}$  such that  $|\mathcal{M}| = n^R$  and  $\lim_{n \to \infty} P_{\text{error}}^n = 0$
- Standard argument, cf. [Cover-Thomas 2006]: Fano's inequality, DPI

$$\begin{aligned} R\log(n) &= H(M|\hat{M}) + I(M;\hat{M}) \\ &\leq 1 + P_{\text{error}}^n R\log(n) + I(M;Y_1^n) \end{aligned}$$

- Consider the Markov chain  $M \to X_1^n \to Z_1^n \to Y_1^n \to S_n \triangleq \sum_{i=1}^n Y_i$ , and a sequence of encoder-decoder pairs  $\{(f_n, g_n)\}_{n \in \mathbb{N}}$  such that  $|\mathcal{M}| = n^R$  and  $\lim_{n \to \infty} P_{\text{error}}^n = 0$
- Standard argument, cf. [Cover-Thomas 2006]: sufficiency

$$R \log(n) = H(M|\hat{M}) + I(M; \hat{M})$$
  
$$\leq 1 + P_{\text{error}}^n R \log(n) + I(M; Y_1^n)$$
  
$$= 1 + P_{\text{error}}^n R \log(n) + I(M; S_n)$$

- Consider the Markov chain  $M \to X_1^n \to Z_1^n \to Y_1^n \to S_n \triangleq \sum_{i=1}^n Y_i$ , and a sequence of encoder-decoder pairs  $\{(f_n, g_n)\}_{n \in \mathbb{N}}$  such that  $|\mathcal{M}| = n^R$  and  $\lim_{n \to \infty} P_{\text{error}}^n = 0$
- Standard argument, cf. [Cover-Thomas 2006]: DPI

$$\begin{aligned} \mathsf{R}\log(n) &= H(M|\hat{M}) + I(M;\hat{M}) \\ &\leq 1 + P_{\text{error}}^n R \log(n) + I(M;Y_1^n) \\ &= 1 + P_{\text{error}}^n R \log(n) + I(M;S_n) \\ &\leq 1 + P_{\text{error}}^n R \log(n) + I(X_1^n;S_n) \end{aligned}$$

- Consider the Markov chain  $M \to X_1^n \to Z_1^n \to Y_1^n \to S_n \triangleq \sum_{i=1}^n Y_i$ , and a sequence of encoder-decoder pairs  $\{(f_n, g_n)\}_{n \in \mathbb{N}}$  such that  $|\mathcal{M}| = n^R$  and  $\lim_{n \to \infty} P_{\text{error}}^n = 0$
- Standard argument, cf. [Cover-Thomas 2006]:

$$R \log(n) = H(M|\hat{M}) + I(M; \hat{M})$$
  

$$\leq 1 + P_{\text{error}}^n R \log(n) + I(M; Y_1^n)$$
  

$$= 1 + P_{\text{error}}^n R \log(n) + I(M; S_n)$$
  

$$\leq 1 + P_{\text{error}}^n R \log(n) + I(X_1^n; S_n)$$

Divide by log(n)

$$R \leq \frac{1}{\log(n)} + P_{\text{error}}^n R + \frac{I(X_1^n; S_n)}{\log(n)}$$

- Consider the Markov chain  $M \to X_1^n \to Z_1^n \to Y_1^n \to S_n \triangleq \sum_{i=1}^n Y_i$ , and a sequence of encoder-decoder pairs  $\{(f_n, g_n)\}_{n \in \mathbb{N}}$  such that  $|\mathcal{M}| = n^R$  and  $\lim_{n \to \infty} P_{\text{error}}^n = 0$
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$$R \log(n) = H(M|\hat{M}) + I(M; \hat{M})$$
  

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$$\leq 1 + P_{\text{error}}^n R \log(n) + I(X_1^n; S_n)$$

• Divide by  $\log(n)$  and let  $n \to \infty$ :

$$R \leq \lim_{n \to \infty} \frac{I(X_1^n; S_n)}{\log(n)}$$

**Upper bound on**  $I(X_1^n; S_n)$ :

 $I(X_1^n;S_n) = H(S_n) - H(S_n|X_1^n)$ 

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Since 
$$S_n \in \{0, ..., n\}$$
,  
 $I(X_1^n; S_n) = H(S_n) - H(S_n | X_1^n)$   
 $\leq \log(n+1) - \sum_{x_1^n \in \{0,1\}^n} P_{X_1^n}(x_1^n) H(S_n | X_1^n = x_1^n)$ 

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Given 
$$X_1^n = x_1^n$$
 with  $\sum_{i=1}^n x_i = k$ ,  $S_n = bin(k, 1-p) + bin(n-k, p)$ :  
 $I(X_1^n; S_n) = H(S_n) - H(S_n | X_1^n)$   
 $\leq log(n+1) - \sum_{x_1^n \in \{0,1\}^n} P_{X_1^n}(x_1^n) H(bin(k, 1-p) + bin(n-k, p))$ 

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Using Problem 2.14 in [Cover-Thomas 2006],  $I(X_1^n; S_n) = H(S_n) - H(S_n | X_1^n)$   $\leq \log(n+1) - \sum_{X_1^n \in \{0,1\}^n} P_{X_1^n}(x_1^n) H(\min(k, 1-p) + \min(n-k, p))$   $\leq \log(n+1) - \sum_{X_1^n \in \{0,1\}^n} P_{X_1^n}(x_1^n) H\left(\min\left(\frac{n}{2}, p\right)\right)$ 

Approximate binomial entropy using CLT, cf. [Adell-Lekuona-Yu 2010]:

$$\begin{split} I(X_1^n;S_n) &= H(S_n) - H(S_n|X_1^n) \\ &\leq \log(n+1) - \sum_{x_1^n \in \{0,1\}^n} P_{X_1^n}(x_1^n) \, H(\min(k,1-p) + \min(n-k,p)) \\ &\leq \log(n+1) - \sum_{x_1^n \in \{0,1\}^n} P_{X_1^n}(x_1^n) \, H\left(\min\left(\frac{n}{2},p\right)\right) \\ &= \log(n+1) - \sum_{x_1^n \in \{0,1\}^n} P_{X_1^n}(x_1^n) \left(\frac{1}{2}\log(\pi e p(1-p)n) + O\left(\frac{1}{n}\right)\right) \end{split}$$

Upper bound on 
$$I(X_1^n; S_n)$$
:  
 $I(X_1^n; S_n) = H(S_n) - H(S_n | X_1^n)$   
 $\leq \log(n+1) - \sum_{x_1^n \in \{0,1\}^n} P_{X_1^n}(x_1^n) H(\min(k, 1-p) + \min(n-k, p))$   
 $\leq \log(n+1) - \sum_{x_1^n \in \{0,1\}^n} P_{X_1^n}(x_1^n) H\left(\min\left(\frac{n}{2}, p\right)\right)$   
 $= \log(n+1) - \frac{1}{2}\log(\pi ep(1-p)n) + O\left(\frac{1}{n}\right)$ 

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Upper bound on 
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 $= \log(n+1) - \frac{1}{2}\log(\pi ep(1-p)n) + O\left(\frac{1}{n}\right)$   
Hence, we have:

$$R \leq \lim_{n \to \infty} \frac{I(X_1^n; S_n)}{\log(n)} = \frac{1}{2}$$

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Upper bound on 
$$I(X_1^n; S_n)$$
:  
 $I(X_1^n; S_n) = H(S_n) - H(S_n | X_1^n)$   
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Hence, we have:

$$R \leq \lim_{n \to \infty} \frac{I(X_1^n; S_n)}{\log(n)} = \frac{1}{2}$$

#### Theorem (Converse)

$$C_{\text{perm}}(\text{BSC}(p)) \leq \frac{1}{2}$$

A. Makur (MIT)



#### 2 Achievability

#### 3 Converse



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#### Theorem (Pemutation Channel Capacity of BSC)

$$C_{\text{perm}}(\text{BSC}(p)) = \begin{cases} 1, & \text{for } p = 0, 1\\ \frac{1}{2}, & \text{for } p \in (0, \frac{1}{2}) \cup (\frac{1}{2}, 1)\\ 0, & \text{for } p = \frac{1}{2} \end{cases}$$



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#### Remarks:

• C<sub>perm</sub>(·) is discontinuous and non-convex

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#### Remarks:

- C<sub>perm</sub>(·) is discontinuous and non-convex
- C<sub>perm</sub>(·) is generally agnostic to parameters of channel

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#### Theorem (Pemutation Channel Capacity of BSC)

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#### Remarks:

- C<sub>perm</sub>(·) is discontinuous and non-convex
- C<sub>perm</sub>(·) is generally agnostic to parameters of channel
- Computationally tractable coding scheme in proof

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#### Theorem (Pemutation Channel Capacity of BSC)

$$C_{\text{perm}}(\text{BSC}(p)) = \begin{cases} 1, & \text{for } p = 0, 1\\ \frac{1}{2}, & \text{for } p \in \left(0, \frac{1}{2}\right) \cup \left(\frac{1}{2}, 1\right)\\ 0, & \text{for } p = \frac{1}{2} \end{cases}$$



#### Remarks:

- C<sub>perm</sub>(·) is discontinuous and non-convex
- C<sub>perm</sub>(·) is generally agnostic to parameters of channel
- Computationally tractable coding scheme in proof
- Proof technique yields more general results

# Thank You!

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