## On Estimation of Modal Decompositions

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## Outline

(1) Introduction

- A Brief History of Modal Decompositions
- Formal Definitions
- Motivation: Embedding of Categorical Data into Euclidean Space
(2) Characterization of Operators
(3) Sample Complexity Analysis
(4) Conclusion


## A Brief History of Modal Decompositions

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- Can we extend these techniques to categorical data?


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- Strong data processing inequalities and related directions: $\chi^{2}$-divergence [Sar58], KL divergence [AG76], and recent work on hypercontractivity [AGKN13], contraction coefficients [MZ15], [PW17], [MZ20], functional inequalities [Rag16], estimation theory, security, and privacy $\left[\mathrm{CMM}^{+} 17\right], \ldots$


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- Lancaster distributions: Mehler's decomposition [Meh66], Lancaster decompositions [Lan58], [Lan69], orthogonal polynomials [Eag64], [Gri69], [Kou96], [Kou98], and recent work [AZ12], [MZ17], ...


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- Non-parametric regression: Alternating conditional expectations (ACE) algorithm [BF85], [Buj85], feature extraction [MKHZ15], [HMZW17], [HMWZ19]


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- Hilbert spaces:

Input space: $\mathcal{L}^{2}\left(X, P_{X}\right) \triangleq\left\{f: X \rightarrow \mathbb{R} \mid \mathbb{E}\left[f(X)^{2}\right]<+\infty\right\}$ with inner product:

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\forall f_{1}, f_{2} \in \mathcal{L}^{2}\left(X, P_{X}\right), \quad\left\langle f_{1}, f_{2}\right\rangle_{P_{X}} \triangleq \mathbb{E}\left[f_{1}(X) f_{2}(X)\right]=\sum_{x \in X} P_{X}(x) f_{1}(x) f_{2}(x)
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and induced $\mathcal{L}^{2}$-norm:

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Output space: $\mathcal{L}^{2}\left(y, P_{Y}\right) \triangleq\left\{g: y \rightarrow \mathbb{R} \mid \mathbb{E}\left[g(Y)^{2}\right]<+\infty\right\}$

## Formal Definitions: Two Equivalent Representations of $P_{X, Y}$

## Definition (Conditional Expectation Operator)

$$
\mathbf{P}_{X \mid Y}: \mathcal{L}^{2}\left(X, P_{X}\right) \rightarrow \mathcal{L}^{2}\left(y, P_{Y}\right) \text { maps any } f \in \mathcal{L}^{2}\left(X, P_{X}\right) \text { to } \mathbf{P}_{X \mid Y} f \in \mathcal{L}^{2}\left(y, P_{Y}\right):
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$$
\forall y \in y, \quad\left(\mathbf{P}_{X \mid Y} f\right)(y) \triangleq \mathbb{E}[f(X) \mid Y=y]
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## Definition (Divergence Transition Matrix)

The divergence transition matrix (DTM), denoted $\mathbf{B} \in \mathbb{R}^{|y| x|x|}$, has $(y, x)$ th entry given by:

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\forall x \in X, \forall y \in y, \quad B(x, y) \triangleq \frac{P_{X, Y}(x, y)}{\sqrt{P_{X}(x) P_{Y}(y)}}
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Remark: DTMs parallel symmetric normalized Laplacian matrices.

## Formal Definitions: SVDs and Modal Decompositions

- $K=\min \{|X|,|y|\}$
- SVD of Conditional Expectation Operator:

$$
\forall i \in\{0, \ldots, K-1\}, \quad \mathbf{P}_{X \mid Y} f_{i}^{*}=\sigma_{i} g_{i}^{*}
$$

- $\sigma_{0} \geq \sigma_{1} \geq \cdots \geq \sigma_{K-1} \geq 0$ are singular values
- $f_{0}^{*}, \ldots, f_{K-1}^{*} \in \mathcal{L}^{2}\left(X, P_{X}\right)$ are orthonormal right singular vectors
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- SVD of DTM:

$$
\mathbf{B}=\sum_{i=0}^{K-1} \sigma_{i} \psi_{i}^{Y}\left(\psi_{i}^{X}\right)^{\mathrm{T}}
$$

- $\psi_{0}^{X}, \ldots, \psi_{K-1}^{X} \in \mathbb{R}^{|X|}$ are orthonormal right singular vectors
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- Relation: $\psi_{i}^{X}(x)=f_{i}^{*}(x) \sqrt{P_{X}(x)}$ for $x \in X$, and $\psi_{i}^{Y}(y)=g_{i}^{*}(y) \sqrt{P_{y}(y)}$ for $y \in y_{\underline{\underline{\beta}}}$


## Formal Definitions: SVDs and Modal Decompositions

## Proposition (SVD Structure)

- Operator Norm: $\sigma_{0}=1, f_{0}^{*}(x)=1$ for all $x \in X$, and $g_{0}^{*}(y)=1$ for all $y \in y$.


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- Maximal Correlations [Hir35, Geb41, Sar58, Rén59]: Using Courant-Fischer-Weyl,

$$
\forall i \in\{1, \ldots, K-1\}, \quad \sigma_{i}=\max _{f, g} \mathbb{E}[f(X) g(Y)]=\mathbb{E}\left[f_{i}^{*}(X) g_{i}^{*}(Y)\right]
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where the maximization is over all $f \in \mathcal{L}^{2}\left(X, P_{X}\right)$ and $g \in \mathcal{L}^{2}\left(y, P_{Y}\right)$ such that $\mathbb{E}\left[f(X)^{2}\right]=\mathbb{E}\left[g(Y)^{2}\right]=1$ and $\mathbb{E}\left[f(X) f_{j}^{*}(X)\right]=\mathbb{E}\left[g(Y) g_{j}^{*}(Y)\right]=0$ for all $j<i$.

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## Proposition (Modal Decomposition of Bivariate Distribution [Hir35, Lan58, Ben73])

$$
\forall x \in X, \forall y \in y, \quad P_{X, Y}(x, y)=P_{X}(x) P_{Y}(y)\left(1+\sum_{i=1}^{K-1} \sigma_{i} f_{i}^{*}(x) g_{i}^{*}(y)\right)
$$

## Motivation: Embedding of Categorical Data into Euclidean Space

Suppose we have:

$$
\begin{aligned}
& x=\{2, \ldots\} \\
& y=\{I S I T, \text { Allerton, ICASSP, ICML, } . .\}
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Dimensionality Reduction:
$|y|$ is large!
Reduce dimension of embedding.

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Want: Low-dimensional embedding of $X$ into Euclidean space $\mathbb{R}^{k}$.

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Modal Decomposition Embedding:

$$
P_{Y \mid X=x}=P_{Y}+\sum_{i=1}^{K-1} \sigma_{i} f_{i}^{*}(x)\left(g_{i}^{*} \cdot P_{Y}\right)
$$

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Modal Decomposition Embedding: When $\sigma_{k+1}$ is small,

$$
\zeta_{k}: X \rightarrow \mathbb{R}^{k}, \quad \zeta_{k}(x)=\left[\sigma_{1} f_{1}^{*}(x) \cdots \sigma_{k} f_{k}^{*}(x)\right]^{T}
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Diffusion Distance Preservation (cf. Laplacian eigenmaps [BN01], diffusion maps [CL06]):

$$
D_{\mathrm{diff}}^{2}\left(P_{Y \mid X=x}, P_{Y \mid X=x^{\prime}}\right) \triangleq \sum_{y \in y} \frac{\left(P_{Y \mid X}(y \mid x)-P_{Y \mid X}\left(y \mid x^{\prime}\right)\right)^{2}}{P_{Y}(y)}=\left\|\zeta_{K-1}(x)-\zeta_{K-1}\left(x^{\prime}\right)\right\|_{2}^{2}
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& \approx\left\|\zeta_{k}(x)-\zeta_{k}\left(x^{\prime}\right)\right\|_{2}^{2} \quad(\text { dimensionality reduction when } k \ll K)
\end{aligned}
$$

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- How do we characterize or identify DTMs?
- Why do we use DTMs or conditional expectation operators to represent $P_{X, Y}$ (instead of, e.g., information density [HV93])? Known relation to mutual $\chi^{2}$-information, ...
- If true distribution $P_{X, Y}$ is unknown but we have training data, how well can we learn $\sigma_{1}, \ldots, \sigma_{k}$ and $\left(f_{1}^{*}, g_{1}^{*}\right), \ldots,\left(f_{k}^{*}, g_{k}^{*}\right)$ ?


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- Representation of Conditional Expectation Operators
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4 Conclusion

## Characterization of DTMs

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- $\mathcal{P}_{0}^{x \times y}=$ relative interior of $\mathcal{P}^{x \times y}$


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## Theorem (Characterization of DTMs)

- $\mathbf{M} \in \mathbb{R}^{|y| \times|X|}$ is a DTM corresponding to a distribution in $\mathcal{P}_{0}^{X \times y}$ if and only if $\mathbf{M}>\mathbf{0}$ (entry-wise) and the spectral norm $\|\mathbf{M}\|_{\text {s }}=1$ :

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$$

- More generally, we have:

$$
\begin{aligned}
\mathbf{B}\left(\mathcal{P}^{X \times y}\right)= & \left\{\mathbf{M} \in \mathbb{R}^{|y| \times|x|}: \mathbf{M} \geq \mathbf{0},\|\mathbf{M}\|_{\mathrm{s}}=1, \exists \psi^{X}>\mathbf{0}, \mathbf{M}^{\mathrm{T}} \mathbf{M} \psi^{X}=\psi^{X},\right. \text { and } \\
& \left.\exists \psi^{Y}>\mathbf{0}, \mathbf{M}^{\mathrm{T}} \boldsymbol{\psi}^{Y}=\psi^{Y}\right\} .
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- DTM function is bijective and continuous.


## Characterization of DTMs

- $\mathcal{P}^{x \times y}=\{$ bivariate distributions over $X \times y$ with strictly positive marginals $\}$
- $\mathcal{P}_{0}^{X \times y}=$ relative interior of $\mathcal{P}^{X \times y}$
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- Proofs utilize Perron-Frobenius theorem.


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## Theorem (Weak Contraction)

- $\min _{Q_{X}>0}\left\|\mathbf{P}_{X \mid Y}\right\|_{Q_{X} \rightarrow P_{Y}}=\left\|\mathbf{P}_{X \mid Y}\right\|_{P_{X} \rightarrow P_{Y}}=1$.
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- Proof uses explicit calculation of adjoint operator $\mathbf{P}_{X \mid Y}^{*}$.


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- Answer: $Q_{X}^{*}=P_{X}$ is unique input Hilbert space that makes $\mathbf{P}_{X \mid Y}$ a weak contraction.


## Outline

(1) Introduction
(2) Characterization of Operators
(3) Sample Complexity Analysis

- Estimation of Dominant Maximal Correlations
- Estimation of Dominant Feature Functions

4 Conclusion

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\forall x \in X, \forall y \in y, \quad \hat{P}_{x, Y}^{n}(x, y) \triangleq \frac{1}{n} \sum_{i=1}^{n} \mathbf{1}_{x_{i}=x} \mathbf{1}_{Y_{i}=y} .
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$$

- Singular value estimates: $\hat{\sigma}_{1} \geq \cdots \geq \hat{\sigma}_{K} \geq 0$
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- Algorithm: Compute SVD of empirical quasi-DTM using, e.g., orthogonal iteration method, QR iteration algorithm (or ACE algorithm), Krylov subspace methods, etc.


## Estimation of $k \in\{1, \ldots, K-1\}$ Dominant Maximal Correlations

- Estimate $\sigma_{1}, \ldots, \sigma_{k}$ using $\hat{\sigma}_{1}, \ldots, \hat{\sigma}_{k}$ under (squared) $\ell^{1}$-norm loss.


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## Corollary (Squared $\ell^{1}$-Risk Bound)

$$
\forall n \geq 16 \log (4 k n), \quad \mathbb{E}\left[\left(\sum_{i=1}^{k}\left|\hat{\sigma}_{i}-\sigma_{i}\right|\right)^{2}\right] \leq \frac{6 k+8 k \log (n k)}{p_{0}^{2} n}
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- Estimate $f_{1}^{*}, \ldots, f_{k}^{*}$ using $\check{f}_{1}^{*}, \ldots, \check{f}_{k}^{*}$ under loss function:

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- First term equals $\sigma_{1}^{2}+\cdots+\sigma_{k}^{2}$ ("rank $k$ approximation" of mutual $\chi^{2}$-information).


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- To estimate $f_{1}^{*}, \ldots, f_{k}^{*}$ within fixed error and confidence, $n$ must be quadratic in $k$.
- Proof exploits: 1) matrix generalization of Bernstein's inequality, and 2) singular value stability result known as Weyl inequality.


## Estimation of $k \in\{1, \ldots, K-1\}$ Dominant Feature Functions

- Estimate $f_{1}^{*}, \ldots, f_{k}^{*}$ using $\check{f}_{1}^{*}, \ldots, \check{f}_{k}^{*}$ under loss function:

$$
\sum_{i=1}^{k}\left\|\mathbf{P}_{X \mid Y} f_{i}^{*}\right\|_{P_{Y}}^{2}-\sum_{i=1}^{k}\left\|\mathbf{P}_{X \mid Y} \check{f}_{i}^{*}\right\|_{P_{Y}}^{2} \geq 0
$$

- First term equals $\sigma_{1}^{2}+\cdots+\sigma_{k}^{2}$.


## Theorem (Sample Complexity Tail Bound II)

$$
\forall 0 \leq \delta \leq 4 k, \quad \mathbb{P}\left(\sum_{i=1}^{k}\left\|\mathbf{P}_{X \mid Y} f_{i}^{*}\right\|_{P_{Y}}^{2}-\left\|\mathbf{P}_{X \mid Y} \breve{f}_{i}^{*}\right\|_{P_{Y}}^{2} \geq \delta\right) \leq(|X|+|y|) \exp \left(-\frac{n p_{0} \delta^{2}}{64 k^{2}}\right)
$$

- To estimate $f_{1}^{*}, \ldots, f_{k}^{*}$ within fixed error and confidence, $n$ must be quadratic in $k$.
- Proof exploits: 1) matrix generalization of Bernstein's inequality, and 2) singular value stability result known as Weyl inequality.
- Analogous results hold for estimation of $g_{1}^{*}, \ldots, g_{k}^{*}$ using $\check{g}_{1}^{*}, \ldots, \check{g}_{k}^{*}$.


## Estimation of $k \in\{1, \ldots, K-1\}$ Dominant Feature Functions

Theorem (Sample Complexity Tail Bound II)

$$
\forall 0 \leq \delta \leq 4 k, \quad \mathbb{P}\left(\sum_{i=1}^{k}\left\|\mathbf{P}_{X \mid Y} f_{i}^{*}\right\|_{P_{Y}}^{2}-\left\|\mathbf{P}_{X \mid Y} \check{f}_{i}^{*}\right\|_{P_{Y}}^{2} \geq \delta\right) \leq(|X|+|y|) \exp \left(-\frac{n p_{0} \delta^{2}}{64 k^{2}}\right)
$$

## Corollary (Mean Squared Error Risk Bound)

For every sufficiently large $n$ such that $\frac{p_{0} n}{64} \geq \frac{1}{|X|+|y|}$ and $\frac{p_{0} n}{4} \geq \log \left(\frac{p_{0} n}{64}(|X|+|y|)\right)$,

$$
\mathbb{E}\left[\left(\sum_{i=1}^{k}\left\|\mathbf{P}_{X \mid Y} f_{i}^{*}\right\|_{P_{Y}}^{2}-\left\|\mathbf{P}_{X \mid Y} \check{f}_{i}^{*}\right\|_{P_{Y}}^{2}\right)^{2}\right] \leq \frac{64 k^{2}\left(\log \left(p_{0} n(|X|+|y|)\right)-3\right)}{p_{0} n}
$$

## Outline

(1) Introduction
(2) Characterization of Operators
(3) Sample Complexity Analysis

4 Conclusion

## Conclusion

## Main Contributions:

- DTMs are entry-wise strictly positive matrices with spectral norm 1.
- Unique Hilbert spaces yield conditional expectation operators that are weak contractions.
- Sample complexity bounds for learning modal decompositions from training data.


## Conclusion

## Main Contributions:

- DTMs are entry-wise strictly positive matrices with spectral norm 1.
- Unique Hilbert spaces yield conditional expectation operators that are weak contractions.
- Sample complexity bounds for learning modal decompositions from training data.


## Main Future Direction:

- Sharpen and generalize sample complexity results using matrix estimation ideas.


## Thank You!

